# UCLA Department of Economics Second-Year Field Examination in Econometrics September 2007

## Instructions:

Solve all three Parts I-III.
Use a separate bluebook for each part.
You have 4 hours to complete the exam.
Calculators and other electronic devices are not allowed.

GOOD LUCK!!!

## PART I (Based on Kyriazidou's course)

### Problem 1:

(a) Suppose that

$$y_{it} = 1 \{ x_{it}\beta + \alpha_i - \varepsilon_{it} \ge 0 \}$$
  $i = 1, ..., N; t = 1, 2$ 

where  $\varepsilon_{it}$  are unobservable variables distributed independently and identically over time conditional on  $(x_{i1}, x_{i2}, \alpha_i)$ ,  $x_{it}$  is a  $1 \times k$  vector of observable variables, and  $\alpha_i$  is an unobservable individual-specific effect. Discuss identification and estimation of  $\beta$ . Assume that cross-sectional sampling is random. Make sure to mention other important assumptions that are used in the identification and consistent estimation of  $\beta$ .

(b) Now suppose that

$$y_{it} = 1 \{ \beta y_{it-1} + \alpha_i - \varepsilon_{it} \ge 0 \}$$
  $i = 1, ..., N; t = 1, 2, 3$ 

where  $\varepsilon_{it}$  are distributed independently and identically over time conditional on  $\alpha_i$  and  $\alpha_i$  is an unobservable individual-specific effect. Assume that  $y_{i0}$  is observed for each i although it is not necessarily generated by the same model as the subsequent  $y_{it}$ 's. Discuss identification and estimation of  $\beta$ . Assume that cross sectional sampling is random.

### Problem 2:

(a) Describe how you would perform Chamberlain's strict exogeneity test in a linear static panel data model of the form

$$y_{it} = x_{it}\beta + \alpha_i + \varepsilon_{it}$$

Make sure to explain the intuition/idea behind the test, to describe the underlying assumptions and to derive its asymptotic distribution. For simplicity you may assume that  $x_{it}$  is scalar.

(b) Discuss how you would perform the same test for the static panel data logit model of the form

$$y_{it} = 1 \{ x_{it}\beta + \alpha_i + \varepsilon_{it} \ge 0 \}$$

(HINT: The strict exogeneity concept need to be strengthened from linear projection to conditional mean independence.)

## PART II (Based on Winkelmann's course)

## Problem 1: Duration analysis

a) Consider a non-negative random variable T (a duration). Assume that the hazard rate is a step function

$$\lambda(t) = \lambda_1 \text{ for } 0 < t < \tau_1$$
  
 $\lambda(t) = \lambda_2 \text{ for } t \ge \tau_1$ 

Derive the survivor function of this model.

- b) Find the expected value of T.
- c) What is the survivor function if the hazard rate is instead

$$\lambda(t) = \alpha t^{\alpha - 1} \lambda_1 \text{ for } 0 < t < \tau_1$$
  
$$\lambda(t) = \alpha t^{\alpha - 1} \lambda_2 \text{ for } t \ge \tau_1$$
?

- d) Write down the log likelihood function for model in c), for a random sample of n possibly right censored observations, where  $d_i$  is an indicator for censoring ( $d_i = 1$  if censored,  $d_i = 0$  else).
- e) Determine the likelihood function of the competing risk model with two independent destinations, where  $\lambda_1(t)$  is the hazard rate of exit to destination 1 and  $\lambda_2(t)$  is the hazard rate of exit to destination 2.

# Problem 2: Count data models

- a) Write down the probability function of a Poisson model with hurdle-atzero.
- b) What are the marginal effects at the extensive and intensive margins, respectively?
- c) In what kind of empirical situation would you consider using the Poisson hurdle model rather than the simple Poisson model?
- d) How can one introduce unobserved heterogeneity into the Poisson hurdle model?
- e) An alternative generalization of the Poisson model is the "zero-inflated" Poisson model. How can one test the hurdle Poisson against the zero-inflated Poisson model?

### PART III (Based on Guggenberger's course)

### Problem 1:

True/Questionable/False? No points are given for just stating true or false, it is the explanation what counts.

- 1) A stationary AR(1) process  $y_t = \rho y_{t-1} + u_t$  (with  $u_t$  iid normal) is ergodic.
- 2) For the Geweke Porter–Hudak estimator  $\hat{d}$  of the long memory parameter d the number of frequencies m used in the pseudo OLS regression has to satisfy  $m/n \to 0$  (where n is the sample size) to ensure that the bias of  $\hat{d}$  goes to 0.
- 3) If in the regression  $y_t = \beta x_t + u_t$ ,  $y_t$  and  $x_t$  are stationary, then there cannot be a spurious regression problem and the OLS estimator of  $\beta$  is consistent.

#### Problem 2:

1) In the linear iid IV model

$$y_i = x_i \theta + u_i,$$
  

$$x_i = z_i \pi + v_i,$$
(1)

derive the asymptotic distribution of the t statistic (that tests  $H_0: \theta = 0$  versus a two sided alternative) under weak instrument asymptotics  $\pi = n^{-1/2}c$  for a fixed constant c. Assume that both  $\theta$  and  $\pi$  are scalars and that the errors are conditionally homoskedastic.

2) Explain the intuition and implementation of Moreira's (Ecta, 2003) conditional likelihood ratio test for the test  $H_0: \theta = 0$  versus a two sided alternative in model (1).

# Problem 3:

The goal is to test  $H_0: \theta = \theta_0$  against  $H_1: \theta > \theta_0$  for given iid data  $X_1, ..., X_n$  with  $EX_i = \theta$ . The test for  $H_0$  is to reject if  $T_n = (\widehat{\theta} - \theta_0)/s(\widehat{\theta}) > c$ , where c is picked so that the type I error is  $\alpha$  ( $\widehat{\theta}$  is a root-n consistent estimator of  $\theta$  and  $s(\widehat{\theta})$  is a consistent estimator of the standard deviation of  $\widehat{\theta}$ ). Compare the following two approaches to do so:

- 1) Using the non-parametric bootstrap, you generate B bootstrap samples, calculate  $\widehat{\theta}^*$ ,  $s(\widehat{\theta}^*)$  for each resample and then calculate  $T_n^* := (\widehat{\theta}^* \theta_0)/s(\widehat{\theta}^*)$ , B times. Let  $q_n^*(1-\alpha)$  denote the  $100(1-\alpha)\%$  quantile of the empirical distribution of  $T_n^*$ . You reject  $H_0$  if and only if  $T_n > q_n^*(1-\alpha)$ .
- 2) Using subsampling, you calculate  $T_{b,i} = (\hat{\theta}_{b,i} \theta_0)/s(\hat{\theta}_{b,i})$ , where  $\hat{\theta}_{b,i}$  and  $s(\hat{\theta}_{b,i})$  are based on the same formula as  $\hat{\theta}$  and  $s(\hat{\theta})$  but instead of using all the data, they only use the data  $\{X_i, ..., X_{i+b-1}\}$  for i = 1, ..., n-b+1. Here b is the blocksize that satisfies  $b/n \to 0$ . Let  $q_{n,b}(1-\alpha)$  denote the  $100(1-\alpha)\%$  quantile of the empirical distribution of  $T_{b,i}$  for i = 1, ..., n-b+1. You reject  $H_0$  if and only if  $T_n > q_{n,b}(1-\alpha)$ .

Discuss the power properties of the two tests. Are they consistent?

### Problem 4:

Explain briefly how the Dickey and Fuller unit root test works.